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Modeling and Forecasting Inflation in Tanzania using ARIMA Models

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ABSTRACT

This research uses annual time series data on inflation rates in Tanzania from 1966 to 2017, to model and forecast inflation using the Box – Jenkins ARIMA technique. Diagnostic tests indicate that the T series is I (1). The study presents the ARIMA (1, 1, 2) model for predicting inflation in Tanzania. The diagnostic tests further imply that the presented optimal model is actually stable and acceptable for predicting inflation in Tanzania. The results of the study apparently show that inflation in Tanzania is likely to continue on an upwards trajectory in the next decade. The study basically encourages policy makers to make use of tight monetary and fiscal policy measures in order to control inflation in Tanzania.

Key Words: Forecasting, Inflation, Tanzania

JEL Codes: C53, E31, E37, E47

INTRODUCTION

Perhaps no subject covers headlines of today's newspaper, economic forums and governments/state agenda as much as inflation (Mbongo, 2014). Inflation is the sustained increase in the general level of prices and services over time (Blanchard, 2000). Inflation is the persistent increase in the level of consumer prices or persistent decline in the purchasing power of money (Webster, 2000). Inflation can also be defined as a situation where the demand for goods and services exceeds their supply in the economy (Hall, 1982). Inflation arises when the central bank creates an excessive supply of money over its demand (Akinboade *et al*, 2004). With a parallel reaction from the production side, the increase in aggregate demand will force prices to go up (Pindiriri, 2012). In reality inflation means that your money cannot buy as much as it could have bought yesterday (Ngailo *et al*, 2014). The negative effects of inflation are

widely recognized (Fenira, 2014). Inflation is one of the central terms in macroeconomics (Enke & Mehdiyev, 2014) as it harms the stability of the acquisition power of the national currency, affects economic growth because investment projects become riskier, distorts consuming and saving decisions, causes unequal income distribution and also results in difficulties in financial intervention (Hurtado *et al*, 2013). Some economists are of the view that a low and stable inflation rate of 3% has a small cost in the economy (Mankiw, 2008). In fact one digit inflation figures are not bad for growth. However, too low inflation, say below, 0%; is not healthy for any economy.

As the prediction of accurate inflation rates is a key component for setting the country's monetary policy, it is especially important for central banks to obtain precise values (Mcnelis & Mcadam, 2004). To prevent the aforementioned undesirable outcomes of price instability, central banks require proper understanding of the future path of inflation to anchor expectations and ensure policy credibility; the key aspects of an effective monetary policy transmission mechanism (King, 2005). Inflation forecasts and projections are also often at the heart of economic policy decision-making, as is the case for monetary policy, which in most industrialized economies is mandated to maintain price stability over the medium term (Buelens, 2012). Economic agents, private and public alike; monitor closely the evolution of prices in the economy, in order to make decisions that allow them to optimize the use of their resources (Hector & Valle, 2002). Decision-makers hence need to have a view of the likely future path of inflation when taking measures that are necessary to reach their objective (Buelens, 2012).

Controlling inflation within acceptable rates is one of the major macroeconomic policies in Tanzania (Laryea & Sumaila, 2001). Despite that, price stability was made a primary concern of the Bank of Tanzania (BOT), the country is experiencing volatility in inflation rates and is yet to permanently secure the target average of 0 to 5 percent annual rate (Ayubu, 2013) and figure 1 below confirms this. To avoid adjusting policy and models by not using an inflation rate prediction can result in imprecise investment and saving decisions, potentially leading to economic instability (Enke & Mehdiyev, 2014). In this study, we seek to model and forecast inflation in Tanzania using ARIMA models.

LITERATURE REVIEW

Nyoni (2018) studied inflation in Zimbabwe using GARCH models with a data set ranging over the period July 2009 to July 2018 and established that there is evidence of volatility persistence for Zimbabwe's monthly inflation data. In another African study, Nyoni (2018) modeled inflation in Kenya using ARIMA and GARCH models and relied on annual time series data over the period 1960 – 2017 and found out that the ARIMA (2, 2, 1) model, the ARIMA (1, 2, 0) model and the AR (1) – GARCH (1, 1) model are good models that can be used to forecast inflation in Kenya. Nyoni & Nathaniel (2019), based on ARMA, ARIMA and GARCH models; studied inflation in Nigeria using time series data on inflation rates from 1960 to 2016 and found out that the ARMA (1, 0, 2) model is the best model for forecasting inflation rates in Nigeria. In the case of Tanzania, Ngailo *et al* (2014) examined inflation based on time series models with a data set ranging over the period January 1997 to December 2010 and revealed that the GARCH (1, 1) and GARCH (1, 2) models are best fit models suitable for predicting inflation in Tanzania. This study is quite different from previous studies such as Ngailo *et al* (2014) in the sense that

we employ annual data and rely on the Box-Jenkins ARIMA technique for estimation purposes; such an approach has not been used before, in the case of Tanzania.

MATERIALS & METHODS

Box – Jenkins ARIMA Models

One of the methods that are commonly used for forecasting time series data is the Autoregressive Integrated Moving Average (ARIMA) (Box & Jenkins, 1976; Brocwell & Davis, 2002; Chatfield, 2004; Wei, 2006; Cryer & Chan, 2008). For the purpose of forecasting inflation rate in Tanzania, ARIMA models were specified and estimated. If the sequence $\Delta^d T_t$ satisfies an ARMA (p, q) process; then the sequence of T_t also satisfies the ARIMA (p, d, q) process such that:

$$\Delta^d T_t = \sum_{i=1}^p \beta_i \Delta^d T_{t-i} + \sum_{i=1}^q \alpha_i \mu_{t-i} + \mu_t \dots \dots \dots [1]$$

which we can also re – write as:

$$\Delta^d T_t = \sum_{i=1}^p \beta_i \Delta^d L^i T_t + \sum_{i=1}^q \alpha_i L^i \mu_t + \mu_t \dots \dots \dots [2]$$

where Δ is the difference operator, vector $\beta \in \mathbb{R}^p$ and $\alpha \in \mathbb{R}^q$.

The Box – Jenkins Methodology

The first step towards model selection is to difference the series in order to achieve stationarity. Once this process is over, the researcher will then examine the correlogram in order to decide on the appropriate orders of the AR and MA components. It is important to highlight the fact that this procedure (of choosing the AR and MA components) is biased towards the use of personal judgement because there are no clear – cut rules on how to decide on the appropriate AR and MA components. Therefore, experience plays a pivotal role in this regard. The next step is the estimation of the tentative model, after which diagnostic testing shall follow. Diagnostic checking is usually done by generating the set of residuals and testing whether they satisfy the characteristics of a white noise process. If not, there would be need for model re – specification and repetition of the same process; this time from the second stage. The process may go on and on until an appropriate model is identified (Nyoni, 2018).

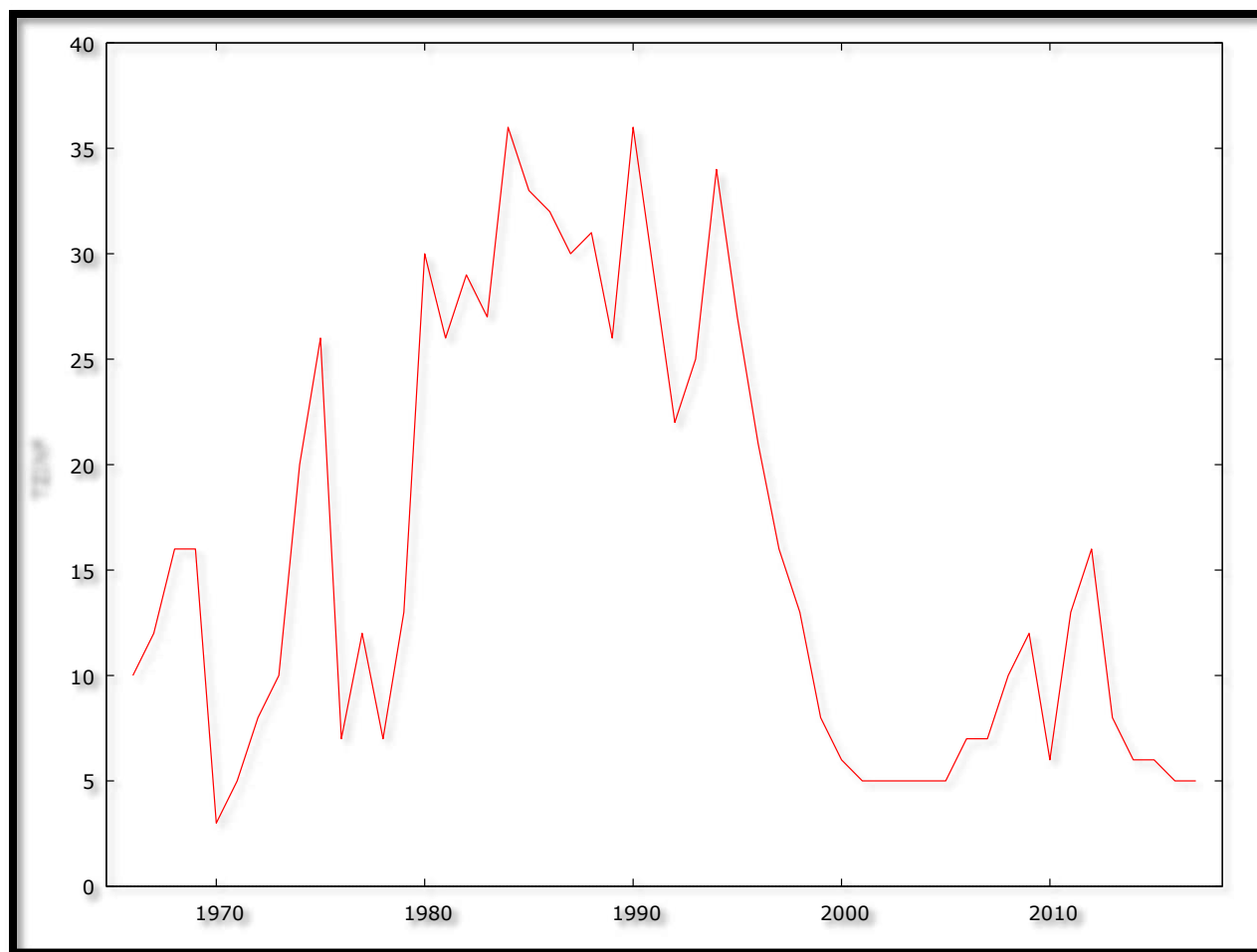
Data Collection

This study is based on a data set of annual rates of inflation in Tanzania (TZINF or simply T) ranging over the period 1966 – 2017. All the data was taken from the World Bank.

Diagnostic Tests & Model Evaluation

Stationarity Tests: Graphical Analysis

Figure 1



The Correlogram in Levels

Autocorrelation function for TZINF ***, **, * indicate significance at the 1%, 5%, 10% levels.

Table 1

| LAG | ACF | PACF | Q-stat. [p-value] |
|-----|------------|------------|-------------------|
| 1 | 0.8202 *** | 0.8202 *** | 37.0385 [0.000] |
| 2 | 0.6911 *** | 0.0561 | 63.8594 [0.000] |
| 3 | 0.6153 *** | 0.1051 | 85.5555 [0.000] |
| 4 | 0.5413 *** | -0.0004 | 102.6987 [0.000] |
| 5 | 0.4898 *** | 0.0525 | 117.0332 [0.000] |
| 6 | 0.4527 *** | 0.0388 | 129.5457 [0.000] |
| 7 | 0.3327 ** | -0.2446 * | 136.4546 [0.000] |

| | | | |
|----|--------|---------|------------------|
| 8 | 0.2123 | -0.1262 | 139.3302 [0.000] |
| 9 | 0.1306 | -0.0382 | 140.4443 [0.000] |
| 10 | 0.0732 | 0.0078 | 140.8024 [0.000] |

The ADF Test in Levels

Table 2: Levels-intercept

| Variable | ADF Statistic | Probability | Critical Values | | Conclusion |
|----------|---------------|-------------|-----------------|-------|----------------|
| T | -2.023015 | 0.2764 | -3.565430 | @ 1% | Non-stationary |
| | | | -2.919952 | @ 5% | Non-stationary |
| | | | -2.597905 | @ 10% | Non-stationary |

Table 3: Levels-trend & intercept

| Variable | ADF Statistic | Probability | Critical Values | | Conclusion |
|----------|---------------|-------------|-----------------|-------|----------------|
| T | -2.321378 | 0.4152 | -4.148465 | @ 1% | Non-stationary |
| | | | -3.500495 | @ 5% | Non-stationary |
| | | | -3.179617 | @ 10% | Non-stationary |

Table 4: without intercept and trend & intercept

| Variable | ADF Statistic | Probability | Critical Values | | Conclusion |
|----------|---------------|-------------|-----------------|-------|----------------|
| T | -1.170820 | 0.2176 | -2.611094 | @ 1% | Non-stationary |
| | | | -1.947381 | @ 5% | Non-stationary |
| | | | -1.612725 | @ 10% | Non-stationary |

Figure 1 and tables 1 – 4 indicate that T is non-stationary in levels.

The Correlogram (at 1st Differences)

Autocorrelation function for d_TZINF ***, **, * indicate significance at the 1%, 5%, 10% levels.

Table 5

| LAG | ACF | PACF | Q-stat. [p-value] |
|-----|----------|---------|-------------------|
| 1 | -0.1468 | -0.1468 | 1.1643 [0.281] |
| 2 | -0.1426 | -0.1677 | 2.2858 [0.319] |
| 3 | -0.0060 | -0.0579 | 2.2878 [0.515] |
| 4 | -0.0973 | -0.1395 | 2.8328 [0.586] |
| 5 | 0.0109 | -0.0442 | 2.8397 [0.725] |
| 6 | 0.2420 * | 0.2113 | 6.3584 [0.384] |
| 7 | -0.0339 | 0.0406 | 6.4289 [0.491] |

8 -0.0498 0.0158 6.5848 [0.582]
 9 -0.0665 -0.0613 6.8693 [0.651]
 10 0.0444 0.0657 6.9993 [0.726]

ADF Test in 1st Differences

Table 6: 1st Difference-intercept

| Variable | ADF Statistic | Probability | Critical Values | | Conclusion |
|----------|---------------|-------------|-----------------|-------|------------|
| T | -8.041948 | 0.0000 | -3.568308 | @ 1% | Stationary |
| | | | -2.921175 | @ 5% | Stationary |
| | | | -2.598551 | @ 10% | Stationary |

Table 7: 1st Difference-trend & intercept

| Variable | ADF Statistic | Probability | Critical Values | | Conclusion |
|----------|---------------|-------------|-----------------|-------|------------|
| T | -8.033070 | 0.0000 | -4.152511 | @ 1% | Stationary |
| | | | -3.502373 | @ 5% | Stationary |
| | | | -3.180699 | @ 10% | Stationary |

Table 8: 1st Difference-without intercept and trend & intercept

| Variable | ADF Statistic | Probability | Critical Values | | Conclusion |
|----------|---------------|-------------|-----------------|-------|------------|
| T | -8.120721 | 0.0000 | -2.612033 | @ 1% | Stationary |
| | | | -1.947520 | @ 5% | Stationary |
| | | | -1.612650 | @ 10% | Stationary |

Tables 5 – 8 show that Z is an I (1) variable.

Evaluation of ARIMA models (without a constant)

Table 9

| Model | AIC | U | ME | MAE | RMSE | MAPE |
|-----------------|----------------|---------|----------|--------|--------|--------|
| ARIMA (1, 1, 1) | 331.1344 | 0.96461 | -0.13782 | 4.328 | 5.86 | 40.042 |
| ARIMA (1, 1, 0) | 330.6554 | 0.95748 | -0.11214 | 4.2795 | 5.95 | 38.308 |
| ARIMA (0, 1, 1) | 330.131 | 0.95891 | -0.1241 | 4.2786 | 5.9181 | 38.917 |
| ARIMA (2, 1, 1) | 332.7998 | 0.97347 | -0.142 | 4.3459 | 5.8398 | 39.952 |
| ARIMA (1, 1, 2) | 332.7944 | 0.9735 | -0.14037 | 4.3578 | 5.8396 | 40.001 |
| ARIMA (2, 1, 2) | 334.7929 | 0.97306 | -0.14016 | 4.3596 | 5.8396 | 40.021 |

A model with a lower AIC value is better than the one with a higher AIC value (Nyoni, 2018). Theil's U must lie between 0 and 1, of which the closer it is to 0, the better the forecast method (Nyoni, 2018). The study will only consider the AIC as the criteria for choosing the best model for forecasting inflation in Tanzania and therefore, the ARIMA (0, 1, 1) model is carefully selected.

Residual & Stability Tests

ADF Tests of the Residuals of the ARIMA (0, 1, 1) Model

Table 10: Levels-intercept

| Variable | ADF Statistic | Probability | Critical Values | | Conclusion |
|----------|---------------|-------------|-----------------|-------|------------|
| R_t | -6.682351 | 0.0000 | -3.568308 | @ 1% | Stationary |
| | | | -2.921175 | @ 5% | Stationary |
| | | | -2.598551 | @ 10% | Stationary |

Table 11: Levels-trend & intercept

| Variable | ADF Statistic | Probability | Critical Values | | Conclusion |
|----------|---------------|-------------|-----------------|-------|------------|
| R_t | -6.702224 | 0.0000 | -4.152511 | @ 1% | Stationary |
| | | | -3.502373 | @ 5% | Stationary |
| | | | -3.180699 | @ 10% | Stationary |

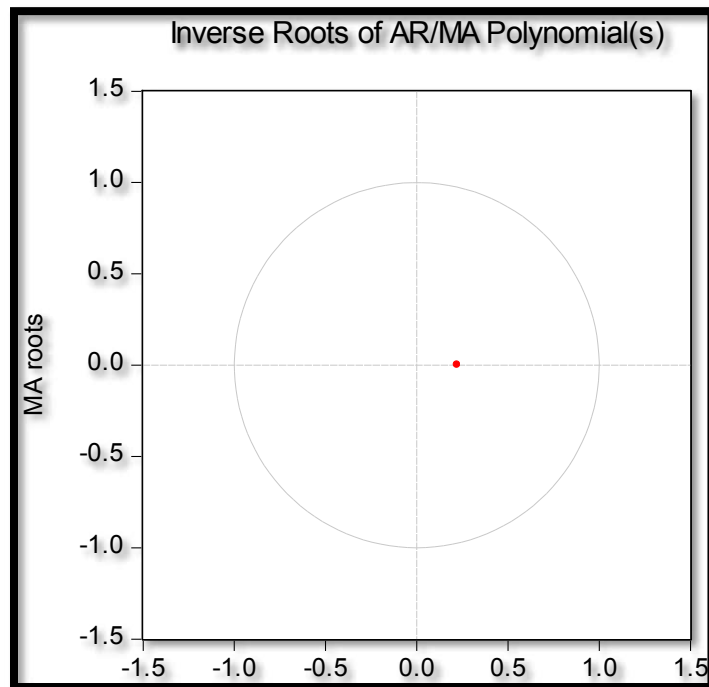
Table 12: without intercept and trend & intercept

| Variable | ADF Statistic | Probability | Critical Values | | Conclusion |
|----------|---------------|-------------|-----------------|-------|------------|
| R_t | -6.746533 | 0.0000 | -2.612033 | @ 1% | Stationary |
| | | | -1.947520 | @ 5% | Stationary |
| | | | -1.612650 | @ 10% | Stationary |

Tables 10, 11 and 12 show that the residuals of the ARIMA (0, 1, 1) model are stationary and hence the ARIMA (0, 1, 1) model is suitable for forecasting inflation in Tanzania.

Stability Test of the ARIMA (0, 1, 1) Model

Figure 2



Since the corresponding inverse roots of the characteristic polynomial lie in the unit circle, it illustrates that the chosen ARIMA (0, 1, 1) model is stable and suitable for predicting inflation in Tanzania over the period under study.

FINDINGS

Descriptive Statistics

Table 13

| Description | Statistic |
|--------------------|-----------|
| Mean | 15.923 |
| Median | 12.5 |
| Minimum | 3 |
| Maximum | 36 |
| Standard deviation | 10.441 |
| Skewness | 0.52114 |
| Excess kurtosis | -1.2087 |

As shown above, the mean is positive, i.e. 15.923%. The minimum is 3% and the maximum is 36%. The skewness is 0.52114 and the most striking characteristic is that it is positive, indicating that the inflation series is positively skewed and non-symmetric. Excess kurtosis was found to be -1.2087; implying that the inflation series is not normally distributed.

Results Presentation¹

Table 14

| ARIMA (1, 1, 2) Model: | | | | |
|---|-------------|----------------|--------|---------|
| $\Delta T_{t-1} = -0.215204\mu_{t-1} \dots \dots \dots [3]$ | | | | |
| P: (0.116) | | | | |
| S. E: (0.1369) | | | | |
| Variable | Coefficient | Standard Error | z | p-value |
| MA (1) | -0.215204 | 0.136917 | -1.572 | 0.116 |

Predicted Annual Inflation in Tanzania

Table 15

| Year | Prediction | Std. Error | 95% Confidence Interval |
|------|------------|------------|-------------------------|
| 2018 | 5.05 | 5.918 | -6.54 - 16.65 |
| 2019 | 5.05 | 7.523 | -9.69 - 19.80 |
| 2020 | 5.05 | 8.841 | -12.27 - 22.38 |

¹ The *, ** and *** means significant at 10%, 5% and 1% levels of significance; respectively.

| | | | | |
|------|------|--------|----------|-------|
| 2021 | 5.05 | 9.986 | -14.52 - | 24.63 |
| 2022 | 5.05 | 11.013 | -16.53 - | 26.64 |
| 2023 | 5.05 | 11.953 | -18.37 - | 28.48 |
| 2024 | 5.05 | 12.823 | -20.08 - | 30.19 |
| 2025 | 5.05 | 13.638 | -21.68 - | 31.78 |
| 2026 | 5.05 | 14.407 | -23.18 - | 33.29 |
| 2027 | 5.05 | 15.137 | -24.61 - | 34.72 |

Table 15 (with a forecast range from 2018 – 2027), clearly show that annual inflation rate in Tanzania is expected to hover around 5.05% within the next decade. However, our 95% confidence interval indicates that inflation in Tanzania is capable of shooting to as high as 34.72% per annum by 2027, ceteris paribus. Therefore, there is need for good macroeconomic policy formulation and implementation in order to maintain price stability in Tanzania.

POLICY IMPLICATION & CONCLUSION

After applying the Box-Jenkins technique, the ARIMA framework was engaged to investigate annual inflation rates in Tanzania over the study period. The study mostly planned to forecast inflation in Tanzania for the upcoming period from 2018 to 2027 and the best fitting model was selected based on how well the model captures the stochastic variation in the data. The ARIMA (1, 1, 2) model is not only stable but also the most suitable model to forecast inflation for the next ten years. In general, inflation in Tanzania; is likely to be hovering around 5.05% per annum over the forecasted period. Based on the results, policy makers in Tanzania should continue to engage proper economic and monetary policies in order to fight such increase in inflation as reflected in the forecasts. In this regard, the BOT is encouraged to rely more on contractionary monetary policy, which should be complimented by a tight fiscal policy.

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